

ISSN 2256-0394 (online) ISSN 2256-0386 (print) 2018, 32, 149–159 doi: 10.2478/eb-2018-0012 https://www.degruyter.com/view/j/eb

HEDGING EFFECTIVENESS FOR INTERNATIONAL INDEX FUTURES MARKETS

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Abstract. This paper investigates the hedging effectiveness of the International Index Futures Markets using daily settlement prices for the period 4 January 2010 to 31 December 2015. Standard OLS regressions, Error Correction Model (ECM), as well as Autoregressive Distributed Lag (ARDL) cointegration model are employed to estimate corresponding hedge ratios that can be employed in risk management. The analyzed sample consists of daily closing market rates of the stock market indexes of the USA and the European futures contracts. The findings indicate that the time varying hedge ratios, if estimated through the ARDL model, are more efficient than the fixed hedge ratios in terms of minimizing the risk. Additionally, there is evidence that the comparative advantage of advanced econometric approaches compared to conventional models is enhanced further for capital markets within peripheral EU countries.

Keywords: Minimum Variance, Hedge Ratio, Effectiveness

JEL Classification: G11, G13

INTRODUCTION

Financial markets have been highly volatile and highly complex in recent decades. As a result, the determination of optimal hedge ratios has emerged as the main subject of discussion for the academic community, but also the subject of monitoring by the majority of financial institutions, investors and businesses. The problem that arises relates to the number of futures that the investor could hold for each underlying unit in order to protect its portfolio against any undesirable market movements. The main objective of the paper is the direct comparison, through a trading strategy process, of the forecasting ability of several econometric approaches that account for the hedging effectiveness. Hedging through trading futures is a common process which is used to control or even reduce the risk of adverse price movements. According to Pennings and Meulenberg (1997), a decisive factor in explaining the success of futures contracts is their hedging effectiveness.

The purpose of this study is to determine whether the parameters that determine the share distribution pattern are significantly influenced by the information available on the market and, consequently, to determine the hedge ratio more 2018/32

effectively. The work calculates the hedge ratios with three alternative econometric methods: Ordinary Least Square Model (OLS), Error Correction Model (ECM) and Autoregressive Distributed Lag cointegration model. It also assesses the effectiveness of alternative risk hedging methods by comparing the risk-benefit ratio by minimal variance, ie minimizing portfolio fluctuations. The hedge ratio is calculated by the fraction of the number of units traded in the futures market over the number of units traded in the spot market. The investor's objectives determine the required hedging strategies. The empirical findings of the paper suggest that more advanced econometric approaches could potentially enhance the effectiveness of the hedging process.

1. LITERATURE REVIEW

One of the most important theoretical issues in Risk Management is finding the optimal risk hedge ratio. The question is to find the optimal method so that investors are protected against possible undesirable market movements at the lowest possible cost. In the literature, the first model was developed by Johnson (1960) and then improved by Ederington (1979) who minimized the variation of the total portfolio using the Ordinary Least Square Method (OLS). However, Park and Bera (1987), as well as Herbst, Kare and Caples (1989) questioned about the importance since it does not take into account the heteroskedasticity that exists between the underlying title and the futures contract. Criticism has also been extended to the fact that the hedge ratios of an OLS model are static, that is, they do not take into account the variability in time-bound variance as well as the alteration of other parameters such as curvature and symmetry. Then, for more effective estimation of hedge ratios, more complex models were used, such as ARCH, GARCH, (Baillie & Myers, 1991; Cecchetti & Cumby, 1988) and cointegration models (Chou, Denis & Lee, 1996; Ghosh, 1993; Lien & Luo, 1993). Baillie & Myers (1991) compare GARCH type with time hedge ratios with hedge ratios and find that hedge ratios are heavier in relation to hedge ratios. These studies ignore the investment profile, which should be taken into account. Ghosh (1993) finds that the minimum variance hedge ratio is underestimated as a result of the wrong model when the combined relationship of futures and underlying securities is not taken into account, and hence error correction is not included in regression. Lien and Luo (1994) argue that cointegration should be taken into account when comparing the ex-post effectiveness of the various hedging strategies.

On the other hand, Lien and Luo (1993) minimized the mean extended Gini coefficient (MEG) in order to find the optimal hedge ratio. Shalit (1995) proved that the hedge ratio which becomes from the MEG approach is identical to the hedge ratio that becomes from the mean-variance approach, when the common distribution of spot and futures returns is normal. Lien and Tse (1998, 2000) and Chen, Lee and Shrestha (2001) in order to estimate the optimal hedge ratio, minimized the generalized semi variance (GSV).

Lien and Tse (1998) in their study found that the GSV hedge ratio is the same as the mean-variance hedge ratio, when the following criteria co-exist: if spot and future prices are normally distributed and if the future prices follow a martingale 2018/32

process. Chen, Lee and Shrestha (2001) developed the GSV hedge ratio to the Mean-GSV hedge ratio. Switzer and El-Khoury (2007) proved that when non symmetry of extreme volatilities is taken in consideration, the hedging performance is improved. A year earlier, Norden (2006) studied the significantly increase of hedging efficiency after the futures split.

Chou, Denis and Lee (1996) used data from the Japan's Nikkei Stock index futures by taking in account different time intervals. Through this method he managed to estimate and compare the hedge ratios between the conventional model and the EC model. Lee, Wang and Chen (2009) used data from international markets like Korean, Singapore, Japan, Taiwan, Hong Kong and United States in order to examine four static and one dynamic hedging model. Their aim was to find the optimal hedge ratio.

During the Asian financial crisis and post-crisis Wang and Hsu (2010) studied the hedge ratio stability of the index futures contracts of Hong Kong, Japan, and Korean. Furthermore, Sah and Pandey (2011) used three models in order to estimate the effective hedge ratio and its hedging effectiveness. His daily data were taken from S&P CNX Nifty futures. Juhl, Kawaller and Koch (2012) used a simple regression on price changes and an error correction model (ECM) in order to find hedge effectiveness. His empirical study proved that when the prices of the hedged item and the hedging instrument are cointegrated, both specifications yield similar results, which depend on the hedge horizon.

Finally, Zhou (2016) extended the methodology in hedging performance by using most effective hedging performance of Real Estate Investment Trusts futures in order to examine the hedge ratios.

2. DATA AND METHODOLOGY

2.1. Data

Our analysis is based on data from international capital markets that refer to developed countries such as US and several core and peripheral EU economies such as Germany, Spain, Italy and Greece. This would enable us to detect possible inefficiencies throughout the comparative analysis. Daily data for the S&P-500, DAX, FTSE/ATHEX-20, IBEX and PSI cash and futures prices are used from 4 January 2010 to 31 December 2015. When a holiday occurs on Thursday, Wednesday's observation is used in its place. Futures prices are always those of the nearby contract because it is the most liquid and active contract. To avoid thin markets and expiration effects (when futures contracts approach their settlement days, their trading volume decreases sharply), we rollover to the next nearest contract one week before the nearby contract expires. For the estimation of hedge ratios we used the period 2010-2014.

The international financial crisis, which was produced in the USA and expanded internationally, did not leave the euro countries intact. The European debt crisis, which intensified in the period 2010-2015, hit most of the southern countries, whose stock markets had a negative course against the German stock market and America. Table 1 shows that the markets with the highest capitalization, i.e. the

DAX index and the S&P index, are a positive instrument while the other markets present a negative medium that does not differ greatly from zero. FTSE / ATHEX-20 index is the worst performing, with the exception of its negative trend, the highest Standard Deviation. This is also due to the country's inability to borrow from the Money Market because of excessive government debt.

Indexes/Futures	Number of Observations	Mean	Standard Deviation	Skewness	Kurtosis
DAX index Spot	1488	0.000473	0.013193	-0.116881	5.109322
Future	1488	0.000473	0.013036	-0.155840	5.300520
S&P-500 index Spot	1488	0.000454	0.010123	-0.317931	7.087520
Future	1488	0.000456	0.010290	-0.333898	7.613981
IBEX index Spot	1467	-0.000045	0.015698	0.426613	8.974296
Future	1467	-0.000043	0.015840	0.413493	9.319258
PSI index Spot	1467	-0.000235	0.013655	0.023012	6.371580
Future	1467	-0.000223	0.014460	0.290209	10.38443
FTSE/ATHEX- 20 index Spot	1467	-0.000852	0.027591	0.239800	6.866535
Future	1467	-0.000803	0.029254	0.550955	10.69623

Table 1. Summary Statistics

2.2. Methodology

The main idea of hedging is to construct a portfolio using investments from the spot market and futures market which will drastically reduce its value fluctuations.

So, we will use a portfolio which has C_s units of a long spot position and C_f units of a short futures position. By S_t and F_t we will denote the spot and the futures prices at time t, respectively. The presence of futures contracts in the portfolio aims to reduce the fluctuations that a portfolio consisting of equities could experience. This portfolio called "hedged portfolio".

We denoted by ΔV_h the price change of the constructed hedged portfolio:

$$\Delta V_h = C_s \Delta S_t - C_f \Delta F_t = C_s \left(\Delta S_t - h \Delta F_t \right)$$

where

 $\Delta S_t = S_{t+1} - S_t$ and $\Delta F_t = F_{t+1} - F_t$ are the price changes of spot and futures positions. Hedge ratio is the fraction of the short futures position over the long spot

position, i.e. $h = C_f/C_s$. Hedging has as main objective to find and determine the optimal hedge ratio h.

Johnson (1960) derives the minimum-variance hedge ratio focusing in the portfolio risk. The measure of the risk in a portfolio is related with the variance of the price change of the hedged portfolio:

$$Var(\Delta V_h) = C_s^2 \left[Var(\Delta S_t) + h^2 Var(\Delta F_t) - 2h Cov(\Delta S_t, \Delta F_t) \right]$$

The minimum-variance hedge ratio is given by

$$h^* = \frac{\text{Cov}(\Delta S_t, \Delta F_t)}{\text{Var}(\Delta F_t)} = \frac{\sigma_{sf}}{\sigma_f^2} = \rho \frac{\sigma_s}{\sigma_f}$$

where

 ρ is the correlation coefficient between ΔS_t and ΔF_t , and σ_s and σ_f are standard deviations of ΔS_t and ΔF_t , respectively.

The estimation of the optimum hedge ratio is implemented applying several econometric models (Lee *et al.*, 2009). Three models are applied in this paper that refers to conventional and also more advanced approaches:

- i) The ordinary least squares (OLS) model,
- ii) The error correction model (ECM) and
- iii) The Autoregressive Distributed Lag (ARDL) cointegration model.

The general econometric model representing the relationship between spot prices and derivatives is the bivariate model of Pesaran (1997):

Assume that S_t and F_t are the spot price and futures price. Then, $\Delta S_t = S_t - S_{t-1}$ and $\Delta F_t = F_t - F_{t-1}$ respectively represent the changes of the prices of spot and futures positions. By Δ we denote the difference operator.

$$\Delta S_{t} = \alpha_{s} (1 - \rho_{s}) - (1 - \rho_{s}) S_{t-1} + \gamma F_{t-1} + \varepsilon_{s,t}$$

$$\Delta F_t = \alpha_f (1 - \rho_f) - (1 - \rho_f) F_{t-1} + \lambda S_{t-1} + \varepsilon_{f,t}$$

and assume that

$$\binom{\varepsilon_{s,t}}{\varepsilon_{f,t}} \sim iid(0,\Sigma), \qquad \Sigma = \begin{pmatrix} \sigma_s^2 & \sigma_{sf} \\ \sigma_{sf} & \sigma_f^2 \end{pmatrix}$$

where

 σ_{sf} is the covariance between ε_s and ε_f , and σ_s^2 and σ_f^2 are variances of ε_s and ε_f , respectively.

In order to estimate the minimum variance hedge ratio we will apply ordinary least squares (OLS) technic, i.e. regression of price changes of the spot on the price changes of the futures. We assume that the spot and futures prices follow a pure random walk and assuming that $\rho_s = \rho_f = 1$, $\gamma = \lambda = 0$ we have:

______ 2018/32

$$\Delta S_t = \varepsilon_{s,t}$$
$$\Delta F_t = \varepsilon_{f,t}$$

If we assume that ε_s and ε_f are jointly normally distributed, we conclude

$$\varepsilon_{s,t} = (\sigma_{sf} / \sigma_f^2) \varepsilon_{f,t} + v_t \tag{1}$$

where

 σ_{sf}/σ_f^2 represents the population coefficient of the regression of $\varepsilon_{s,t}$ on $\varepsilon_{f,t}$, and v_t is distributed independently of $\varepsilon_{f,t}$. Then we could present the ordinary least squares model as follows:

$$\Delta S_t = \alpha + \beta \Delta F_t + v_t$$

(with $\beta = \sigma_{sf} / \sigma_f^2$ under the jointly normality condition). The estimation of β will give the estimation of minimum variance hedge ratio, h^* ,

Alternatively, according to the long-run equilibrium relationship between spot and futures prices we apply the error correction model. The condition for arbitrage connects spot and futures prices. Following this, it is natural to suppose that they cannot drift far apart in the long run. Therefore, if each series follows a random walk process, then we expect this two series to be cointegrated. We take the assumption that the spot price S_t and futures price F_t are unit-root processes and at the same time are co-integrated. Assuming that

$$\mid \rho_s \mid <1, \ \rho_f =1, \ \gamma \neq 0, \ \lambda =0 \text{ or } \rho_s =1, \ \mid \rho_f \mid <1, \ \gamma =0, \ \lambda \neq 0 \text{ we have:}$$

$$\Delta S_{t} = \alpha_{s} (1 - \rho_{s}) - (1 - \rho_{s}) S_{t-1} + \gamma F_{t-1} + \varepsilon_{s,t}$$
(2)

and

$$\Delta F_t = \varepsilon_{f,t}$$

We write $\varepsilon_{s,t} = \beta \varepsilon_{f,t} + v_t$ and use this expression to rewrite the previous equations as:

$$\Delta S_{t} = \alpha + \theta z_{t-1} + \beta \Delta F_{t} + v_{t}$$

where

$$\alpha = \alpha_s (1 - \rho_s)$$
, $\theta = -(1 - \rho_s)$, $\beta = (\sigma_{sf} / \sigma_f^2)$ and $z_{t-1} = S_{t-1} - \gamma / (1 - \rho_s) F_{t-1}$ is referred to as the error-correction term at time t .

We will apply the least squares method to the model. The estimation of β corresponds to the optimal hedge ratio h*. If the spot price and futures price series are found to be cointegrated, then the hedge ratio can be estimated in two steps (Ghosh, 1993). The first step involves the estimation of the following cointegrating regression:

$$S_t = a + bF_t + e_t \tag{3}$$

The second step involves the estimation of the following error correction model:

$$\Delta S_t = \mu + \lambda u_{t-1} + \beta \Delta F_t + v_t \tag{4}$$

where

 u_t is the residual series from the cointegrating regression in Equation (3). The estimation of β from (4) will give the estimation of the hedge ratio.

The Autoregressive Distributed Lag cointegration model studied first by Pesaran (1997). He stated that the existence of a long-run relationship between the spot price S_t and futures price F_t does not depend on whether F_t is I(1) (i.e. $\rho_f = 1$).

If we suppose that there exists a single long-run relationship between S_t and F_t , then it have to be $|\rho_s| < 1$, $\gamma \neq 0$, $\lambda = 0$ or $|\rho_f| < 1$, $\lambda \neq 0$, $\gamma = 0$. We assume that $|\rho_s| < 1$, $\gamma \neq 0$, $\lambda = 0$ (without loss of generality). Then we have that:

$$\Delta F_t = \alpha_f (1 - \rho_f) - (1 - \rho_f) F_{t-1} + \varepsilon_{f,t}$$

Since we supposed that ε_s and ε_f are jointly normally distributed (1), we have from (2):

$$\Delta S_t = \alpha_1 + \alpha_2 S_{t-1} + \alpha_3 F_{t-1} + \beta \Delta F_t + V_t \tag{5}$$

where

$$\begin{split} &\alpha_1 = \alpha_s \, (1 - \rho_s) - \alpha_f \, (\sigma_{sf} \big/ \sigma_f^2) (1 - \rho_f), \ \alpha_2 = -(1 - \rho_s), \\ &\alpha_3 = [\gamma + (\sigma_{sf} \big/ \sigma_f^2) (1 - \rho_f)] \text{ and} \end{split}$$

 $\beta = (\sigma_{sf}/\sigma_f^2)$. At this point, we could mention that Equation (5) is a single-step process. This means that it is not the same from the two-step method used by the Equation (4).

3. RESULTS

Table 2 presents the results of the estimate hedge ratio based on the methodology described above.

Table 2. Estimated Hedge Ratios

Index		Per Day
DAX index	$oldsymbol{eta}^{ extit{oLS}}$	1.004252
	$oldsymbol{eta}^{ECM}$	1.015479
	$oldsymbol{eta}^{ARDL}$	1.015949
S&P 500 index	$oldsymbol{eta}^{ extit{oLS}}$	0.958953
	$oldsymbol{eta}^{ECM}$	0.956242
	$oldsymbol{eta}^{ARDL}$	0.963976

Index		Per Day
FTSE/ATHEX-20 index	$eta^{ extit{OLS}}$	0.881013
	$oldsymbol{eta}^{ECM}$	0.868288
	$oldsymbol{eta}^{ARDL}$	0.896831
PSI index	eta^{oLS}	0.888204
	$eta^{ extit{ iny ECM}}$	0.882088
	$oldsymbol{eta}^{ARDL}$	0.899438
IBEX index	eta^{oLS}	0.980981
	$eta^{ extit{ iny ECM}}$	0.985129
	$oldsymbol{eta}^{ARDL}$	0.985811

We note that the hedge ratio is lower than the unit for all markets except for the DAX index, which is close to 1. Moreover, the hedge ratio in all methods is similar except Greece's index showing some variations. With regard to the risk hedging strategy for an administrator, the ARDL model can provide the most effective risk hedging.

3.1. Unit root test

To make sure that the time series data is stationary, we used the augmented Dickey-Fuller test (Dickey & Fuller, 1981) and PP test (Phillips & Perron, 1988). Both tests appeared in Table 3 and indicated that the first order differences of the spot and the futures prices in the five markets generate stationary series at the 1% significance level.

Table 3. Results of the Unit-Root Tests on Spot and Futures Prices

		ADF	PP
Dax index	Spot	-40,8868**	-41,1653**
	Futures	-26,8623**	-42,2563**
S&P index	Spot	-31,4564**	-43,4556**
	Futures	-45,2455**	-45,3669**
FTSE-20 index	Spot	-18,2566**	-37,2567**
	Futures	-19,5897**	-37,4577**
IBEX index	Spot	-20,1445**	-40,1245**
	Futures	-17,5461**	-43,5778**
PSI index	Spot	-11,2489**	-39,2456**
	Futures	-12,2899**	-40.4877**

^{*}indicates significance at 5% level, ** indicates significance at 1% level.

3.2. Hedge ratios and hedging effectiveness measures

A popular measure of hedging effectiveness was created by Ederington (1979) and since applied in numerous empirical studies, is the percentage reduction in variance:

$$Variance _ \text{Re } duction = \frac{Var(\Delta S_t) - Var(R_{H,t})}{Var(\Delta S_t)}$$

which compares the variance of the returns of the hedged portfolios $Var(R_{H,t})$ to the variance of the unhedged positions $Var(\Delta S_t)$.

Hedging performance, measured using variance reduction, is calculated out-of-sample using data from 01/01/2015-31/12/2015. The results of the variance reduction in the out of sample analysis summarizes in the table 4.

	Ordinary Least Square	Error Correction Model	ARDL Model
DAX index	98,084%	98,037%	98,094%
S&P index	93,524%	93,516%	93,635%
FTSE-20 index	89,766%	89,593%	90,936%
IBEX index	97,172%	97,154%	97,693%
PSI index	94,017%	93,919%	94,476%

Table 4. Hedging effectiveness results

We notice that the largest variance reduction is achieved in the DAX index while the lowest in the FTSE-20 index. This means that for DAX we have higher hedging ratios. The ARDL Model also achieves the largest variance reduction across all stock indices.

CONCLUSION

This paper examines different methods of hedging in American and European stock portfolios with the help of futures contracts. Hedge ratios are calculated using three alternative econometric methods: the Ordinary Least Square Model (OLS), the Error Correction Model (ECM), and the Autoregressive Distributed Lag (ARDL) cointegration model.

It also evaluates the effectiveness of alternative hedging methods based on variance reduction. The results show that *Autoregressive Distributed Lag* cointegration model is more efficient than the fixed hedge ratios in terms of minimizing the risk and this superior performance of the advanced econometric models is pronounced when analyzing data from the periphery capital markets.

2018/32

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